For those 40 years of age, the lifetime risk of developing heart failure (HF) is 20%. Diagnosis is associated with a high level of disability, healthcare costs, and mortality (ie, ≈50% within 5 years of diagnosis). Earlier detection of HF opens the possibility to test means to preserve cardiac function and change the natural history of disease onset. The rapid growth of health systems with electronic health records (EHRs) in the past 15 years opens new opportunities to implement early detection surveillance. However, the success that any health system will have in detecting HF early with predictive methods will depend on patient-level data diversity, quantity, and density. Information that can be gained on populations of patients from longitudinal EHR data can be used to individualize care for a given patient. Access to these data in combination with the rapid evolution of modern machine learning and data mining techniques offers a potentially promising means to accelerate discoveries that can be readily translated to clinical practice.

Rapid growth in the application of machine learning methods to EHR for predicting future patient outcomes is fostering a deeper understanding of best practices and the trade-offs between model performance and patient coverage (ie, the percent of potentially eligible patients for which risk assessment can be completed). Model utility for a given system will strongly depend on the length of the prediction window, the length of the observation window, and the diversity, quantity, and density of the available data. A longer prediction window (ie, the duration of time before diagnosis that disease is detected) would likely increase the potential for successful prevention, but accuracy may decrease. Model performance will improve, to some degree, as the observation window (ie, duration of time before the prediction window from which data are used) increases. But, patient coverage decreases (ie, more
WHAT IS KNOWN
- EHR data can be used to predict events including onset of diseases, but there is little known about the practical tradeoffs between data requirement and model utility.
- Rapid growth in the application of machine learning methods to EHR for predicting future patient outcomes is fostering a deeper understanding of modeling best practices.

WHAT THE STUDY ADDS
- Detailed characterization of the practical tradeoffs between data requirements and model utility is possible and shows the value and limitations of heterogeneous data sources to predictive models in EHR data.
- Guidelines are possible for the minimum amount and type of data needed to train effective disease onset predictive models using longitudinal EHR data.

patients will be excluded from risk assessment) because of the more stringent data requirement. The diversity (Table 1) of available data may also influence how well a model performs. Performance will likely improve as more different data domains are used. However, as the number of required data domains increases, it will become more challenging for health systems to make use of developed predictive models. Moreover, some data domains (eg, diagnostic codes) may be more useful than others, and there may be a diminishing return when the threshold of ≥1 qualifying combinations is reached. Finally, performance will likely depend on the overall amount (ie, the number of case and control patients) and density of the available data.

Relatively little is known about the tradeoffs among these factors. To this end, we used data from a large health system to examine how well machine learning tools perform in relation to the prediction window length, observation window length, and the diversity, quantity, and density of the available EHR data.

Methods
Study Design, Population, Setting, and Source of Data
For this study, a nested case–control design was used on a large patient population from a health system that provided access to longitudinal EHR data. The source data were from the Geisinger Clinic, which is a multispecialty group practice with a large primary care practice. All cases and controls were primary care patients. Geisinger includes 41 outpatient clinics in central and northeastern Pennsylvania. EpicCare EHR® was installed at Geisinger before 2001. The study was approved by an institutional review committee, and the subjects gave informed consent.

Definition and Selection of Cases and Controls
Criteria for incident onset of HF were adopted from Gurwitz et al8 and defined in detail elsewhere,9,10 but relied on qualifying International Classification of Diseases-Ninth Revision (ICD-9) codes for HF with a minimum of 3 clinical encounters occurring within 12 months of each other. The date of diagnosis was assigned to the earliest of the 3 dates. Finally, incident cases had to have at least 18 months before the first occurrence of a HF ICD-9 code diagnosis without an indication that HF had been previously diagnosed or treated. Analysis was limited to patients who were 50 to 84 years of age at the time of HF diagnosis. Applying these criteria, a total of 1684 incident HF cases were identified between 2003 and 2010, where the average primary care patient census was n=240,000, of whom 28% were between 50 and 84 years of age.

Up to 10 eligible primary care clinic-, sex-, and age-matched (5-year age intervals) controls were selected for each incident HF case for a total of 13,525 control patients. Primary care patients were eligible as controls if they had no HF diagnosis before the date: 1 year post-HF diagnosis of the matched case. Control subjects were required to have their first office encounter within 1 year of the incident HF patient’s first office visit and have ≥1 office encounters 30 days before or any time after the case’s HF diagnosis date to ensure similar duration of observations among cases and controls. Nine or ten controls were identified for 49% of the cases. Only 1.5% of the cases had just 1 or 2 controls.

Structured Data Extraction From Longitudinal EHR Patient Data
EHR encounter data were extracted for all cases and controls from each of the 8 data-type domains (Table 1). Encounters included outpatient face-to-face and phone visits and inpatient visits. As shown in Table 1, the number of unique variables in each data type substantially varied.

For the diagnoses data type, ICD-9 codes from outpatient visits and problem lists are treated separately. This means that the same ICD-9 code associated with the different events were treated as separate variables. For the medications data type, normalized drug names (ie, combining all branded names and the generic name for a medication) were used. Dosing information was not used. For the hospitalization data type, the primary and secondary ICD-9 codes associated with the hospital admission were used.11 The laboratories and vital signs data types contain information about not only the test order but also the numeric results of the test. From the cardiovascular imaging order data type only information about the imaging order type (2-dimensional echo, stress echo, etc.), associated ICD-9 codes, and any left ventricular ejection fraction (LVEF) values were used. No other test results of the imaging orders were available.

For the diagnoses, hospitalizations and medication data types, different levels of representations were explored. For the diagnoses and hospitalizations data types, the ungrouped representation was the ICD-9 disease code (401.0, 492.0, 584.90, etc.), whereas the grouped representation was the hierarchical condition categories from the Centers for Medicare & Medicaid Services: for example, hypertension, chronic obstructive pulmonary disease, renal failure.12 For the medications data type, the ungrouped representation was the normalized drug name (hydrochlorothiazide, atenolol, lisinopril, furosemide, etc.), whereas the grouped representation was the pharmacy subclass (thiazide diuretics, β-blockers, angiotensin-converting enzyme inhibitors, loop diuretics, etc.). Grouping reduces the number of available features and creates features that are less sparse. As shown in Table 1, the number of unique features drops from 12,616 to 189 for the diagnoses data type, from 7,071 to 189 for the hospitalizations data type, and from 3,952 to 631 for the medications data type.

Feature Construction From Longitudinal EHR Patient Data
A feature vector representation for each patient was generated based on the patient’s extracted EHR data. The EHR data were represented as a temporal sequence of events (eg, a patient can have multiple diagnoses of hypertension on different dates) in the observation window. Events of the same feature type (eg, antihypertensive medication) within the observation window were represented by ≥1 relevant summary statistics. The summary functions can include simple feature values like Booleans, counts, means, and complex feature values that can take into account temporal information (eg, trend and temporal variation). In this study, basic aggregation functions were applied and included Booleans for categorical variables and means for numeric
An initial set of experiments was performed comparing several different data-driven machine learning classification approaches to select the ones to use in this study. The different classifiers compared were random forest, support vector machine, k nearest neighbor, decision tree, logistic regression, and L1-regularized logistic regression. The L1-regularized logistic regression and random forest models were selected for superior predictive performance, computational efficiency, and because they represent different modeling approaches that would be interesting to compare.

Predictive Modeling

An initial set of experiments was performed comparing several different data-driven machine learning classification approaches to select the ones to use in this study. The different classifiers compared were random forest, support vector machine, k nearest neighbor, decision tree, logistic regression, and L1-regularized logistic regression. The L1-regularized logistic regression and random forest models were selected for superior predictive performance, computational efficiency, and because they represent different modeling approaches that would be interesting to compare.

In this study, the information gain measure was used to select the top 200 most discriminating features from the original ≥3400 features. Interactions between the variables were not considered. Next L1-regularized logistic regression and random forest models were trained and evaluated. Unconditional logistic regression models with adjustment for continuous age were used to account for residual confounding. Internal cross-validation was used to determine the model hyperparameters: the number of features in the logistic regression model and number of trees in the random forest model.

A sequence of experiments was completed to measure how well data-driven machine learning predictive modeling performs in relation to the prediction window length, the observation window length, and the diversity, quantity, and density of the available EHR data. Models were trained and evaluated as detailed below. Fixed prediction window (1 year) and observation window (2 years) were used, except where specified where these parameters were evaluated.

- Data Diversity: Eight different data types (Table 1) were individually assessed on model performance. For the diagnoses and hospitalizations data types, different levels of representations of the ICD-9 codes were explored. For the medications data type, different levels of representation of the medications were explored. Models were also trained and evaluated using data-type combinations starting with demographics/health behaviors and then adding vitals, diagnoses, medications, laboratories, Framingham Heart HF signs and symptoms, hospitalizations and imaging to characterize the impact of using more data types.

- Prediction window length: varied from 30 days to 5 years.

- Observation window length: varied from 60 days to 5 years.

- Data quantity: varied in the amount of randomly selected training data that ranged from 114 to 11,400 patients and evaluated

![Table 1. Electronic Health Records Data Types, Examples, Number of Unique Variables, and Number of Grouped Variables](image)

<table>
<thead>
<tr>
<th>Data Type</th>
<th>Examples</th>
<th>No. of Unique Variables</th>
<th>No. of Grouped Variables</th>
</tr>
</thead>
<tbody>
<tr>
<td>Diagnoses*</td>
<td>ICD-9 codes (from outpatient visits and problem lists treated separately)</td>
<td>12616</td>
<td>189</td>
</tr>
<tr>
<td>Medications†</td>
<td>β-blockers, loop diuretics, etc</td>
<td>3952</td>
<td>631</td>
</tr>
<tr>
<td>Laboratories‡</td>
<td>Cholesterol, glucose, eGFR, etc</td>
<td>2336</td>
<td>...</td>
</tr>
<tr>
<td>Hospitalization§</td>
<td>ICD-9 codes associated with admission</td>
<td>7071</td>
<td>189</td>
</tr>
<tr>
<td>Demographics and health behaviors</td>
<td>Age, sex, race/ethnicity, smoking, and alcohol use</td>
<td>5</td>
<td>...</td>
</tr>
<tr>
<td>Cardiovascular imaging orders</td>
<td>2D echo, transesophageal echo, stress echo, etc</td>
<td>18</td>
<td>...</td>
</tr>
<tr>
<td>Vitals‡</td>
<td>Pulse, systolic blood pressure, diastolic blood pressure, height, weight, and temperature</td>
<td>6</td>
<td>...</td>
</tr>
<tr>
<td>Framingham HF signs and symptoms</td>
<td>Positive and negative mentions of acute pulmonary edema, ankle edema, dyspnea on ordinary exertion, paroxysmal nocturnal dyspnea, etc</td>
<td>28</td>
<td>...</td>
</tr>
</tbody>
</table>

eGFR indicates estimated glomerular filtration rate; HF, heart failure; and ICD-9, International Classification of Diseases-Ninth Revision.

*ICD-9 codes from outpatient visits and problem lists are treated separately.
†Normalized drug names, ignoring the dosing information, were used.
‡The laboratories and vitals data types contain information about the test order and the numeric result of the test.
§Primary and secondary ICD-9 codes associated with the hospital admission were used.
†The imaging order type, associated ICD-9 codes, and any left ventricular ejection fraction values were used.

In this study, the information gain measure was used to select the top 200 most discriminating features from the original ≥3400 features. Interactions between the variables were not considered. Next L1-regularized logistic regression and random forest models were trained and evaluated. Unconditional logistic regression models with adjustment for continuous age were used to account for residual confounding. Internal cross-validation was used to determine the model hyperparameters: the number of features in the logistic regression model and number of trees in the random forest model.

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- Prediction window length: varied from 30 days to 5 years.

- Observation window length: varied from 60 days to 5 years.

- Data quantity: varied in the amount of randomly selected training data that ranged from 114 to 11,400 patients and evaluated

![Figure 1. Relation of prediction window, observation window for use of data, and the index and diagnosis dates for cases and the same relative times for controls.](image)
on a held out testing data set while keeping the case/control ratio constant.
• Data density: different amounts of data were selected based on the number of encounters in the observation window.

Ten-fold cross-validation was used in all modeling iterations (which includes both the initial feature selection step and the classifier training step) to make efficient use of the available data and to obtain an estimate of how well the model can generalize to an independent data set. Prediction performance was assessed by the mean and SE of the area under the receiver operating characteristic curve (AUC) metrics computed over the cross-validation folds. The experiments used a combination of software tools developed in house and based on open source packages for Python and R.

Results

Performance as a Function of the Data Diversity

Figure 2 shows the prediction performance for ungrouped versus grouped representations for the diagnoses, medications, and hospitalizations data types when used individually (ie, unique formats represented in the EHR). The approach to grouping, based on hierarchical condition categories, relies on a knowledge-driven hierarchy and was used to eliminate correlated and redundant features as a means to enhance model performance by allowing other useful features to be selected by the model. For each of these 3 data types (Figure 2), the grouped representation performs as well as or better than the nongrouped representation. The logistic regression models using the grouped representation were also less complex than those using the nongrouped representations as measured by the number of model features. For ICD-9 codes, the difference in the number of selected variables was 43 versus 59 for the diagnoses type and 11 versus 40 for the hospitalizations type. For the medications data type, both representations have 13 selected variables, but the grouped representation has significantly better performance: AUC 0.708 versus 0.674, $P<0.01$. Table 2 lists some examples of the top selected grouped features for the diagnoses, hospitalizations, and medications data types. For the rest of the experiments in this study, we used the grouped representations for the diagnoses, hospitalizations, and medications data types.

Individually, diagnosis and medication order data types performed best (Figure 3A), followed by laboratory test and hospitalization data types. Demographics/health behaviors, imaging test orders, and vitals data types perform the worst on their own.

Figure 3B shows prediction performance as data types are used in combination starting with demographics/health behaviors and adding vitals, diagnoses, medications, laboratories, hospitalizations, and imaging. The combination sequence is based on our assessment of the availability of each of the data types from most to least common. Use of only the demographics/health behaviors and vitals data types did not yield models that perform well. Adding the diagnoses data type results in a large increase in model performance (AUC 0.733 versus 0.581; $P<0.00001$). The addition of the medications data type results in a small but significant performance improvement (AUC 0.752 versus 0.733; $P<0.05$). Adding the laboratories data type did not significantly improve model performance. In contrast, the addition of the hospitalization data type significantly improved performance (AUC 0.791 versus 0.756; $P<0.001$). Finally, the addition of the imaging data types did not significantly improve performance.

Table I in the Data Supplement lists the selected variables and related statistics for the L1-regularized logistic regression models trained on 90% of the data from one of the cross-validation folds using all the data types. The variables selected for the other cross-validation folds were similar. Hospitalization, diagnosis, medication, and laboratory features dominated. A few of the Framingham HF signs/symptoms and imaging

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**Table 2. Example Feature Groups for the Diagnoses, Hospitalizations, and Medications Data Types**

<table>
<thead>
<tr>
<th>Diagnoses</th>
<th>Hospitalizations</th>
<th>Medications</th>
</tr>
</thead>
<tbody>
<tr>
<td>Blood disorders</td>
<td>Cardiac disorders</td>
<td>ACE inhibitors</td>
</tr>
<tr>
<td>Cardiac disorders</td>
<td>ENT disorders</td>
<td>A-β blockers</td>
</tr>
<tr>
<td>Cerebrovascular disease</td>
<td>Eye diseases</td>
<td>Angiotensin II receptor antagonists</td>
</tr>
<tr>
<td>Coronary atherosclerosis</td>
<td>Gastrointestinal disorders</td>
<td>Antiarrhythmics type III</td>
</tr>
<tr>
<td>Diabetes mellitus</td>
<td>Injuries</td>
<td>Antidiabetic combinaisons</td>
</tr>
<tr>
<td>Gastrointestinal disorders</td>
<td>Lung disease</td>
<td>β-blockers cardiodselective</td>
</tr>
<tr>
<td>Heart arrhythmias</td>
<td>Musculoskeletal disorders</td>
<td>Cardiac glycosides</td>
</tr>
<tr>
<td>Heart infection</td>
<td>Skin diseases</td>
<td>Coumarin anticoagulants</td>
</tr>
<tr>
<td>Kidney disorders</td>
<td>Symptom disorders</td>
<td>Insulin</td>
</tr>
<tr>
<td>Lung disease</td>
<td>Loop diuretics</td>
<td></td>
</tr>
<tr>
<td>Metabolic diseases</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Musculoskeletal disorders</td>
<td>Platelet aggregation inhibitors</td>
<td></td>
</tr>
<tr>
<td>Neoplasm</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Neurological diseases</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Polyneuropathy</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Vascular disease</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

ACE indicates angiotensin-converting enzyme; and ENT, ear, nose, throat.
variables were also selected. None of the demographic/health behaviors or vital features was selected.

Table II in the Data Supplement shows the top variables selected in the random forest model based on the variable importance score. A random forest classifier was built, measuring how much the classification output changes as each input feature variable is randomly changed. Important input features are expected to substantially improve classification output value, whereas changes to an unimportant input feature will not affect the output much. The variable importance scores from the random forest classifier trained on each of the 10 cross-validation folds were averaged and represented as composite scores to then rank the features. The random forest classifier uses many more features from all data types except imaging than the logistic regression classifier. For the rest of the experiments in this study, we used all 8 of the data types. Descriptions of all the variables appearing in Tables I and II in the Data Supplement are listed in Table III in the Data Supplement.

Performance as a Function of Prediction Window Length
The impact on model performance was evaluated for a window length of 30 days to 5 years (Figure 4) with a fixed 2-year observation window. The prediction window length specifies the amount of time before the true clinical HF diagnosis that the prediction is performed (Figure 1). As expected, performance for both classifiers declines in relation to increasing prediction window length (ie, the time before the actual clinical diagnosis). Prediction performance is at or above 0.80 AUC for prediction windows ≤ 1 year. Performance is at 0.74 AUC for a 2-year prediction window. Performance then declines rapidly for prediction window lengths longer than 2 years.

Performance as a Function of Observation Window Length
The impact of model performance was evaluated in relation to the observation window length ranging from 60 days to 5 years.
Observation window length specifies the length of time in the patient record from which data are extracted to represent the patient (Figure 1). The longer the observation window, the more data points from the patient record are used except, of course, where patient time under observation is shorter than the observation window. For both classifiers, prediction performance improves steadily as the observation window length increases up until 2 years where the performance reaches 0.79 AUC. Longer observation windows (3, 4, and 5 years) had minimal impact on model performance (up to 0.80 AUC).

Performance as a Function of the Training Set Size
The original data set is randomly split into 25% (3802 patients) for testing and 75% (11407 patients) for training. The training set was randomly sampled (without replacement) and varied from 114 to 11407 patients in size to train the classifiers and then used to score the patients in the test set. Fifty iterations of training and testing for each training set size are performed, and the average and SE of the 50 AUC scores are plotted as a function of the training set size (Figure 6).

Prediction performance improves substantially from an AUC 0.653 to an AUC of 0.786, \( P < 0.00001 \), for the logistic regression classifier, as the size of the training set increases from 114 to 3802 patients. After 3802 patients, model performance begins to asymptote (from 0.786 to 0.796 AUC) as the training set size increases to 11407 patients. Performance of the random forest classifier is slightly better at all data set sizes but has a similar trend.

Performance as a Function of the Training Data Density
An estimate of the data density (ie, in-person and phone encounters) in the observation window (set at 2 years for this analysis) is computed for each case and control by first accounting for each instance of a diagnoses, medications, laboratories, imaging orders, and vitals data types along with the associated date. Next, the total number of distinct dates is computed. This number of encounters is then used as the data density estimate. A total of 18% of cases (Figure 7, red line) and 25% of controls (Figure 7, blue line) had <10 encounters. Half of the controls have <20 encounters and half of the cases have <25 encounters.

To measure prediction performance as a function of the training data density, a series of regularized logistic regression and random forest classifiers were trained and evaluated using 10-fold cross-validation on different subsets of the training data selected using 2 methods: (1) based on increasing data density in the observation window and (2) randomly without replacement to meet a specified data set size. The density-based selection was performed by including case and control patients who have at least the minimum number of specified encounters in the 2-year observation window. The bar graphs in Figure 8 show that as the density threshold increases, the number of selected cases (red bars) and controls (blue bars) that have the minimum number of required encounters in the observation window decreases. Each bar graph is labeled with a percent of the total patients who are cases, where cases comprise 11% of all patients. This percentage changes slightly when both case and control patients are restricted to those with ≥10 encounters in the observation window (ie, 12%–13%) and then increases incrementally to 15% at 20 encounters and finally to 26% at 60 encounters. The solid blue and red curves in Figure 8 plots the mean and SE of the cross-validated AUC scores (vertical axis) as a function of the data set density (horizontal axis) expressed as the minimum number of encounters in the observation window for the regularized logistic regression and random forest classifiers, respectively. Predictive performance improves significantly when data are restricted to those with at least 10 to 15 encounters in the 2-year observation window, a range where the exclusion of cases and controls is similar. Model performance continues to improve when data are restricted to those with more encounters, but the selective exclusion of controls is substantial when compared with cases. As a comparative baseline, models trained using the same amount of data but selected randomly instead of by data density are evaluated. For each threshold value, 50 iterations of training and testing are performed using randomly selected data sets that are the same size as the corresponding density-based data set. The dashed blue and red curves in Figure 8 plots the mean (and SE) of the 50 AUC scores for each threshold data set size for the regularized logistic regression and random
forest classifiers, respectively. Performance holds steady for moderate data set sizes (down to 6000 patients) and then slowly drops off as the data set size decreases. These results seem to indicate the importance of having less but denser data over having more but sparser data.

**Discussion**

The adoption of EHRs in US healthcare is contributing to an explosive growth in diverse longitudinal patient data that will increasingly support population-level analyses to support individualized patient care. Predictive analytics represent one important domain that is relevant to early detection of disease, avoidable care (eg, hospital admission), and clinical decision support, among numerous other applications. The quantitative performance of a predictive model, represented by the AUC, is often a dominant focus of concern, but this represents only one of several factors that will dictate the overall utility of and the likelihood that a model will be adopted for clinical care (Table 3). In this article, we explored how the quantity and diversity of EHR data influences model performance defined by both the AUC and other features. We consider implications for implementation of predictive models specifically for the early detection of HF and then consider issues with generalizability.

The demand for more diverse, voluminous, or higher density of data to improve model performance creates an inevitable tension by reducing the size of the market that can use the model. This tension does not necessarily lead to the creation of a single model that is designed to the lowest common denominator. Rather, the tradeoff influences the family of models that might be created to accommodate differences in the data that are available to specific provider groups. In our analysis, model performance was most strongly influenced by the diversity of data, basic feature construction, and the length of the observation window. In raw form, EHR data are highly diverse, represented by thousands of variants for disease coding, medication orders, laboratory measures, and other data types. It seems
obvious that some level of feature construction that relies on well-established ontologies should improve model performance. We examined this specific issue because it had not been previously explored. Use of ontologies for ICD-9 codes and medication classes each substantially improves model AUC, reduces the number of features, or both. Modest curation of data using existing ontologies substantially reduces data diversity without loss of information and possibly a gain in information that both supports improvement to model performance, parsimony, and coherence. Moreover, this modest data curation also simplifies standardization of data across different provider entities. This finding may suggest that more advanced forms of data curation that are beyond traditional ontologies may further improve model performance and parsimony.

The adoption of predictive models will depend, in part, on the data diversity requirements. Adoption is more favorable when the same or similar model performance can be achieved with fewer types of data. In our analysis, model performance was considerably better when using either diagnoses or medication order data alone when compared with other data types (laboratories, hospitalization, demographics/health behaviors, image orders, and vitals). These 2 domains may be dominant because they also represent the most common sources of data and also offer information relative to a patient’s status over a longer period of time than a clinical measure. This does not necessarily mean that larger quantities of laboratory or vital sign data that could be generated from wearable devices could also improve model performance for those specific domains. Notably, the combination of diagnoses, medications, demographics/health behaviors, and vital sign data does not yield an AUC that is significantly better than diagnoses and medication data alone (ie, AUC of 0.752 versus 0.740; \( P = 0.23 \), data not shown). Although we cannot generalize to predicting other health problems, the findings are sensible given that information about smoking and alcohol use habits and vital signs are likely to be captured by diagnoses and medication data. For example, ICD-9 codes for hypertension combined with related medication orders may be more informative than intermittent measures of blood pressure. One limitation to our assessment of data types is related to the image order data. Other than echo order data on LVEF, we did not extract test results for other imaging orders.

We examined a 12-month prediction window, varying the length of the observation window. Although a longer observation window increases the volume of data, it also reduces the number of patients to whom the model could be applied in a given provider group. Not surprisingly, model performance improved substantially as the length of the observation window increased. Performance began to plateau for observation windows longer than 2 years. For HF, this finding may suggest that signals of early detection in EHR data may only exist within 1 to 3 years of diagnosis. However, in modeling experiments, the observation window was treated as a static data set represented by summary statistics. It is possible that longer observation windows improve model performance when using more advanced models that can make temporal connections among features or events.

The most surprising finding in this study is how sensitive the model performance is to the size of the training set. Substantial improvements to our predictive model were achieved by increasing the training set size from 1000 patient records to 4000 records. Although the training set size may be influenced by other factors (eg, type of data, case/control ratio, and data density), this finding, in particular, may pose the greatest challenge to application of models in clinical practice. It is likely that any predictive model that is developed for use in clinical practice will have to be adapted to the data available in a clinical practice. That is, model development work largely offers a guide on data requirements and data curation, not on the specific formulation that could be applied to patient data. As such, the training set size may be the most influential factor in adoption of predictive models for smaller practice groups. The case–control data set for this study included 1684 HF cases and 13,525 controls from a modest size clinical practice. The incident HF cases accrued over a 7-year period. It is unlikely that most practice groups will have equivalent or greater denominators, represented by the combination of years of data and size of the patient population. Although we do not have evidence of the generalizability of this finding to other areas of predictive modeling, we can conclude that there is likely to be a strong trade-off in model performance and the size of the training set and that smaller practices may benefit from pooling of data.

The quantity and diversity of available EHR data are highly heterogeneous among patients, and this poses potential methodological challenges in using EHR data for predictive modeling purposes. In contrast to epidemiological data that are collected using a structured protocol and time schedule, EHR data accrue in an unscheduled manner and with a modest amount of structure defined by clinical protocols. For example, vital signs are collected during most visits, but often with variation in the protocol. The frequency, type, and quantity of laboratory tests ordered can also be widely inconsistent among practitioners and their patients. Other data such as ICD-9 codes and medication orders are documented using standardized schemes, but physicians vary in how they document and what they order. Patients also vary in how much care they use, independent of disease burden. In our analysis, we found that data density strongly influences model performance. Using a 2-year observation window and 1-year prediction window, model performance substantially improved.
when a minimum of 10 encounters was required, but even at these thresholds ≥20% of patients in a primary care practice would be excluded from risk assessment using the HF model. Performance continues to improve by restricting the data to patients with >15 encounters. In this experiment, the number of encounters between cases and controls are matched. As such, the performance gain cannot be explained by the difference in number of encounters between cases and controls. Some of the improvement may be because of a better representation of the patient derived from more data. Some of the additional gains in AUC may be because of the larger fraction of cases compared with controls being selected with higher density threshold values. More work needs to be done to better understand this behavior. Because some individuals who develop HF do so after a singular acute event, such as after a heart attack or a viral cardiomyopathy, there are likely to be a subset of people whose HF diagnosis follows a much more rapid trajectory and are more likely to have fewer medical encounters in the months to years before their diagnosis.

Several features (ie, α-β-blockers orders, positive JV distension, negative pulmonary edema, loop diuretic orders, and cardiac disorders) in the final model (Table I in the Data Supplement) may suggest a covert HF diagnosis, but this does not necessarily mean that HF is actually present or even that individuals with those features will develop HF. The predictive importance of a feature (Table II in the Data Supplement) takes into account both the discriminability and frequency of the variable. Features that may indicate HF often ranked relatively low in importance (eg, positive JV distension ranked 97th, α-β-blockers ranked 89th, negative pulmonary edema ranked 46th). When these 5 features were eliminated from the model, the AUC was only reduced from 0.796 to 0.785. More importantly, there is considerable noise in EHR data, and it is difficult to state with certitude that an individual patient is at risk of a future diagnosis of HF simply because these features appear in their record. For example, Framingham HF signs and symptoms were common among cases and controls and had relatively little impact on model performance. Predictive models offer the means to discriminate the meaning of thousands of data points that simply cannot be deciphered by manual review. Moreover, even if a patient had HF before the recorded diagnosis date, the patient is not being managed accordingly. The value of predictive models in this context is to offer the means to conduct routine surveillance to surface and more effectively manage these patients before they are hospitalized for the first time with HF.

Our findings are consistent with previous work using EHR data. In predicting mortality over different time horizons from end-stage renal disease, machine learning models that contain all available predictors outperform those that do not. But variable groups varied by utility, where exclusion of any particular group did not lead to a meaningfully reduction in risk prediction. Moreover, prediction performance decreased as the prediction window size increased (from 1 day to 3 years). Separately, renal function disease progression was modeled using machine learning methods with and without temporal information on data features. Prediction performance improved for temporal and nontemporal models as the observation window size increased with peak performance occurring between 2 and 3 years of patient data. In a different study, significant differences in the accuracy of hypoglycemia prediction using machine learning models for patients with type 2 diabetes mellitus were observed as a function of data density (number of self-monitored blood glucose data points) and prediction window size (1–24 hours). Model performance increased with denser data and with smaller prediction window size.

There are many possible directions for future work. First, the approach and methods need to be validated on larger patient data sets from multiple healthcare systems and additional disease targets to better understand the generalizability of the data characteristic impacts on predictive modeling performance. More in-depth analysis of the characteristics and distribution of the top risk factors captured by the predictive model is also needed. In this study, the prediction and observation windows were explored separately; understanding the combination of them is also important. Exploration of novel approaches to data curation that rely on combinations of data elements to characterize variation in disease complexity, disease control, and other features may improve model performance or coherence or both. It will be important to explore data density, training set size, and other requirements in greater depth and to develop methods that reduce the data demands as a means to increase adoption. Finally, use of modeling strategies that account for the sequencing and temporal distance of features in the observation window may offer an approach to increase model performance with the same amount of data.

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None.

References


