Identification of Hospital Outliers in Bleeding Complications After Percutaneous Coronary Intervention

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Background—Post-percutaneous coronary intervention (PCI) bleeding complications are an important quality metric. We sought to characterize site-level variation in post-PCI bleeding and explore the influence of patient and procedural factors on hospital bleeding performance.

Methods and Results—Hospital-level bleeding performance was compared pre- and postadjustment using the newly revised CathPCI Registry® bleeding risk model (c-index, 0.77) among 1292 National Cardiovascular Data Registry® hospitals performing >50 PCIs from 7/2009 to 9/2012 (n=198 4998 procedures). Using random effects models, outlier sites were identified based on 95% confidence intervals around the hospital’s random intercept. Bleeding 72 hours post-PCI was defined as: arterial access site, retroperitoneal, gastrointestinal, or genitourinary bleeding; intracranial hemorrhage; cardiac tamponade; nonbypass surgery–related blood transfusion with preprocedure hemoglobin ≥8 g/dL; or absolute decrease in hemoglobin value ≥3 g/dL with preprocedure hemoglobin ≤16 g/dL. Overall, the median unadjusted post-PCI bleeding rate was 5.2% and varied among hospitals from 2.6% to 10.4% (5th, 95th percentiles). Center-level bleeding variation persisted after case-mix adjustment (2.8%–9.5%; 5th, 95th percentiles). Although hospitals’ observed and risk-adjusted bleeding ranks were correlated (Spearman ρ: 0.88), individual rankings shifted after risk-adjustment (median Δ rank order: ±91.5; interquartile range: 37.0, 185.5). Outlier classification changed postadjustment for 29.3%, 16.1%, and 26.5% of low-, non-, and high-outlier sites, respectively. Hospital use of bleeding avoidance strategies (bivalirudin, radial access, or vascular closure device) was associated with risk-adjusted bleeding rates.

Conclusions—Despite adjustment for patient case-mix, there is wide variation in rates of hospital PCI-related bleeding in the United States. Opportunities may exist for best performers to share practices with other sites. (Circ Cardiovasc Qual Outcomes. 2015:8:00-00. DOI: 10.1161/CIRCOUTCOMES.113.000749)

Key Words: percutaneous coronary intervention ■ quality improvement

Each year, ≈600,000 percutaneous coronary intervention (PCI) procedures are performed in the United States, yet there are few outcomes-based quality indicators of PCI performance. Currently used performance measures include in-hospital PCI mortality and risk-standardized 30-day readmissions after PCI; however, one of the challenges of these quality improvement metrics is whether they can be modified by alterations in care processes and consequently improved upon. Another limitation of in-hospital PCI mortality is that the rates are low, limiting the variation across hospitals, as well as the usefulness of this metric to judge performance.8

Recent attention has focused on PCI-related bleeding as a potential hospital quality indicator. Bleeding is the most common noncardiac complication of PCI and is associated with increased morbidity, mortality, and cost. Because bleeding after PCI has been consistently associated with known patient characteristics, such as older age, female sex, and renal insufficiency, bleeding risk models have been developed and validated to provide accurate estimates of post-PCI bleeding risk and, therefore, guide therapy and improve patient outcomes. PCI-related bleeding risk can be modified by provider factors, such as use of bivalirudin and radial access, and vascular closure devices may potentially reduce bleeding complications in certain populations, but have not been definitively tested. However, data suggest that the use of these approaches (collectively termed bleeding avoidance strategies (BAS)) is variable. Recently, the National Cardiovascular Data Registry CathPCI Registry began including hospital risk-adjusted post-PCI bleeding rates in its provider reports. Furthermore, PCI bleeding has been designated as a quality...
WHAT IS KNOWN

• Bleeding complications after percutaneous coronary intervention are associated with poor prognosis and have been targeted as a site quality metric.

WHAT THE STUDY ADDS

• Significant hospital-level variation in post-percutaneous coronary intervention bleeding exists but was attenuated after adjustment for patient clinical characteristics, demonstrating the importance of risk-adjustment for appropriate site comparisons.
• Substantial variation in hospital post-percutaneous coronary intervention bleeding rates persisted after accounting for patient case-mix.
• Hospital use of bleeding avoidance strategies (bivalirudin, radial access, or vascular closure device) was associated with reduced rates of bleeding.
• Wide implementation of bleeding avoidance strategies and sharing of practices from best-performing sites may be useful for quality improvement initiatives to reduce post-percutaneous coronary intervention bleeding.

Definitions and Outcomes

CathPCI Registry data definitions can be found online at www.ncrdr.com/webncre/cathpci/home/datacollection. The use of BAS was defined as the use of any of the following during PCI: radial artery access, bivalirudin, or vascular closure device. Academic or teaching status was defined as the presence of an internship, residency, or fellowship program at an institution.

Bleeding after PCI was the primary outcome for our study. In 2009, the CathPCI Registry implemented a new data collection form with more detailed data variables related to bleeding events. Using these data elements, the CathPCI Registry postprocedure bleeding definition was recently revised to capture (1) potentially unreported bleeding events using more objective laboratory data and (2) important bleeding complications, such as intracranial hemorrhage and cardiac tamponade, that were not available in prior versions.7 Bleeding was reported by trained site data collectors and defined based on any of the following events occurring within 72 hours of the procedure: arterial access site bleeding, either overt external bleeding or a hematoma >10 cm for femoral access, >5 cm for brachial access, or >2 cm for radial access; retroperitoneal, gastrointestinal, or genitourinary bleeding; intracranial hemorrhage; cardiac tamponade; decrease of ≥3 g/dL in hemoglobin post-PCI in patients with preprocedure hemoglobin ≤16 g/dL; or postprocedure nonbypass surgery–related blood transfusion in patients with a pre-PCI hemoglobin of ≥28 g/dL.

CathPCI Registry Bleeding Risk Adjustment Model

The CathPCI Registry bleeding model was recently updated to include bleeding events not captured in previous CathPCI Registry data collection forms and to be more comparable to bleeding definitions used in other studies.16 This new standard CathPCI Registry bleeding model was developed using data from PCI procedures performed between February 2008 and April 2011. Factors significantly associated with bleeding complications within 72 hours of PCI were identified using multivariable logistic regression. The final model includes the following clinical factors, such as demographic variables (female sex, age, and body mass index), comorbidities (diabetes mellitus, prior congestive heart failure, prior PCI, cerebrovascular disease, peripheral vascular disease, chronic lung disease, and chronic kidney disease), clinical presentation characteristics (ST-segment–elevation myocardial infarction, New York Heart Association heart failure class, ejection fraction, cardiac arrest within 24 hours, cardiogenic shock, baseline hemoglobin), and procedural variables (preprocedure thrombolysis in myocardial infarction flow, number of diseased vessels, Society for Cardiovascular Angiography and Interventions lesion class, lesion segment). The model had good discrimination (c-index: 0.77) and was well calibrated.

Statistical Analysis

Patient, procedure, and hospital characteristics were described across tertiles (low, average, and high) of unadjusted hospital PCI bleeding rates. Categorical variables were presented as frequencies and percentages, and continuous variables were summarized as medians with interquartile ranges. Comparisons among categorical and continuous variables were performed using Pearson χ² and Wilcoxon rank-sum tests, respectively. Unadjusted hospital bleeding rates were calculated with hospital included as a random effect variable, whereas adjusted bleeding rates were determined after accounting for all variables in the CathPCI Registry bleeding model in addition to the hospital. Models incorporating hospital as a random effect allow for formal statistical testing of whether any observed variation in outcome is because of differences among hospitals (variance parameter estimates >0 and with P≤0.05) versus simple sampling variation. Hospitals were classified as outliers using the 95% confidence interval (CI) around the hospital’s random intercept, which is a shrunk estimator representing the log odds of bleeding for each hospital (Data Supplement).25 Hospitals for which the lower 95% CI limit was >1 were considered to have high outlier status; hospitals for which the upper 95% CI limit was <1 were considered to have low outlier status; and hospitals whose 95% CI

WHAT THE STUDY ADDS

metric in the Centers for Medicare and Medicaid Services Acute Care Episode Demonstration program.25

Although there is interest in the adoption of post-PCI bleeding as a site performance measure, evidence to support it has been limited. To date, overall variability in hospital rates of post-PCI bleeding has not been reported, and the influence of patient or procedural factors on hospital bleeding rates has not been examined. Therefore, we sought to (1) characterize hospital-level variation in post-PCI bleeding rates; (2) assess the contribution of patient case-mix to variation in bleeding rates among sites; and (3) explore whether hospital factors, including use of BAS, are associated with post-PCI bleeding.

Methods

The CathPCI Registry is a national quality improvement program jointly sponsored by the American College of Cardiology and the Society for Cardiovascular Angiography and Interventions. This registry provides in-hospital data on patients undergoing cardiac catheterization and PCI from 1,400 hospitals in the United States. Details about the CathPCI Registry have been previously published.26

Study Sample

We included all PCI procedures performed between July 2009 and September 2012 with data reported using version 4 of the CathPCI Registry data collection form. From a total of 20,241,611 index PCI procedures performed at 1,383 participating sites, we excluded patients with missing bleeding data (n=1,097) and sites that reported no bleeding events or did not record any hemoglobin values (n=13,141 patients from 73 sites). We also excluded patients undergoing coronary artery bypass grafting during the index hospitalization (n=24,380 procedures) and sites with <50 PCIs (n=545 procedures from 18 sites). Our final analysis population consisted of 1,984,998 PCI procedures performed at 1,292 sites (Figure 1).
included 1 were considered to be nonoutliers. We assessed the relationship between unadjusted and adjusted bleeding rates by calculating the absolute values of change in rank order for hospitals and by quantifying the association between unadjusted and adjusted hospital rankings using Spearman rank correlation coefficient.

We performed sensitivity analyses. We increased the threshold for site exclusion from <50 PCIs to <150 PCIs to explore the stability of bleeding rate estimates. Although there has been increasing acceptance of random effects modeling to compare hospital outcomes, current methodology used to assess participating CathPCI Registry hospital performance in site reports use of nonrandom effects models (Data Supplement). Therefore, analyses were repeated using these same fixed effect models. We calculated observed bleeding rates for hospitals by dividing the observed number of bleeds by the total number of admissions. The expected number of bleeds for each hospital was determined using the validated CathPCI Registry bleeding model to tabulate the sum of the predicted probabilities of bleeding for each patient at that hospital. Hospital adjusted rates were then obtained by multiplying the ratio of observed to expected number of events (a measure commonly used to assess hospital performance) by the population bleeding rate. Hospital outlier status for observed rates was defined using the 95% CI for a hospital’s observed bleeding rate divided by the population bleeding rate, using the same definitions of outlier status as in the main analysis. Hospital outlier status for adjusted rate was defined similarly using the 95% CI for a hospital’s ratio of observed-to-expected bleeding rates. Standard errors were based on the binomial distribution rather than Poisson because this is also the statistical methodology used by the CathPCI Registry. Finally, we determined Spearman correlation coefficients to explore a limited number of hospital and procedural factors that might be associated with hospital-level post-PCI bleeding. For all analyses, statistical tests were 2-sided, and a P<0.05 was considered statistically significant. Analyses were performed at the Duke Clinical Research Institute using SAS software (version 9.2, SAS Institute, Cary, NC) and STATA (Release 11, StataCorp, College Station, TX). This study was approved by the Duke University Medical Center Institutional Review Board and is determined to qualify for a waiver of informed consent.

Results

Hospital-Level Variation in Bleeding

Among the 1984998 PCI procedures at 1292 sites, the overall median unadjusted rate of post-PCI bleeding was 5.2%. As shown in Figure 2A, there was wide variation in bleeding rates across hospitals. Center-level variation in composite bleeding rates ranged as follows: 5th, 10th, 25th, 75th, 90th, and 95th percentiles were 2.6%, 3.0%, 3.9%, 6.9%, 8.8%, and 10.4%, respectively.

We also assessed the rates of the individual bleeding components that make up this composite (Table 1). We found that the individual bleeding end point rates generally followed the composite results across hospital-level bleeding tertiles. Because the need for red blood cell transfusion remains somewhat subjective, we...
also investigated whether differences in the composite end point were because of differential thresholds in transfusion. We found that hospital pretransfusion hemoglobin values were highly consistent across all 3 bleeding tertiles (Table 1).

**Patient, Procedural, and Hospital Characteristics Among Hospital Bleeding Tertiles**

Patient, procedure, and hospital characteristics according to tertiles of unadjusted hospital-level bleeding (low: <4.41%; average: 4.41%–6.23%; and high: >6.24%) are shown in Tables 2 and 3. Compared with patients in the low and average tertiles, patients treated at hospitals in the high tertile were more often of nonwhite race and more often had a history of prior myocardial infarction and prior congestive heart failure, but were less likely to have undergone prior coronary revascularization.

Patients in the high tertile also more frequently presented with ST-segment—elevation myocardial infarction and heart failure and more frequently underwent PCI for emergency and salvage indications than patients in the lower 2 tertiles. As shown in Table 3, hospitals in the high versus low and average tertiles of hospital-level bleeding were more often teaching hospitals and had lower median annual PCI volumes. Compared with PCI procedures in the lower 2 tertiles, procedures at centers in the high tertile of hospital-level bleeding were more often with unfractionated heparin and glycoprotein IIb/IIIa inhibitors. In contrast, bivalirudin, radial access, and vascular closure devices were used most frequently during PCI procedures performed at hospitals in the low tertile of post-PCI bleeding, as reflected in the highest proportion of hospital use of any BAS in this tertile.

**Effect of Risk-Adjustment on Hospital Bleeding Performance Rates and Ranks**

After applying the CathPCI Registry risk model to adjust for case-mix, variation remained in hospital bleeding rates (Figure 2B). Hospital risk-adjusted bleeding rates ranged with 5th, 10th, 25th, 50th, 75th, 90th, and 95th percentiles of 2.8%, 3.2%, 4.1%, 5.3%, 6.7%, 8.3%, and 9.4%, respectively. We examined the effect of risk-adjustment on hospital rank and outlier status. Figure 3 depicts the relationship between hospital rank based on unadjusted versus risk-adjusted rates of bleeding. In general, the values were well correlated (Spearman correlation coefficient: 0.88). However, individual hospital rankings among the 1292 sites shifted with risk-adjustment, the median change in rank order observed was ±91.5 (interquartile range: 37.0, 185.5). Next, hospital outlier status was determined based on either unadjusted or adjusted bleeding rates (Table 4). Before adjustment for case-mix, 300 sites were classified as low outliers (lower than expected bleeding rates), and 370 sites were considered high outliers (higher than expected bleeding rates). Overall, risk-adjustment shifted outlier status for 22.1% (n=286) of sites, with 29.3% (n=88) of low outlier, 16.1% (n=100) of nonoutlier, and 26.5% (n=98) of high outlier sites changing classification (Table 4).

**Factors Associated With Variation in Hospital-level Bleeding**

In addition to patient case-mix, we explored hospital and procedural factors that might be associated with hospital-level post-PCI bleeding. Of 1292 total sites, 39.6% (n=511) were teaching hospitals. The median bleeding rates at academic versus nonacademic hospitals were similar (5.4% and 5.1%, respectively). Overall, the median annual hospital PCI volume was 391.2 cases (interquartile range: 203.7, 656.8). Hospital annual PCI volume was not correlated with center-level risk-adjusted bleeding (Spearman correlation coefficient: 0.02). In contrast, bleeding rates were associated with several procedural factors. On average, hospitals used heparin and glycoprotein IIb/IIIa inhibitors in 52.4% and 31.4% of PCI cases, respectively. Increasing hospital use of each medication was correlated with greater bleeding.
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(Spearman ρ: 0.27 for heparin and 0.40 for glycoprotein IIb/IIIa inhibitor, P<0.0001 for both). The average hospital-level percent-age use of any BAS was 79.1%. We found a correlation between higher hospital percentage use of BAS and lower risk-adjusted bleeding (Spearman correlation coefficient: −0.26).

Sensitivity Analyses

To ensure stability of hospital bleeding estimates, we performed a sensitivity analysis after raising the threshold for exclusion from this study from <50 to <150 PCIs per site. Using this criterion, 60 sites were excluded. Among the remaining 1236 hospitals, the 5th and 95th percentiles for bleeding rates were 2.6% and 10.3%, respectively, which were similar to the rates of 2.6% and 10.4% from the original analysis.

Table 2. Patient and Procedure Characteristics by Tertile of Hospital-Level Unadjusted Bleeding Rates

<table>
<thead>
<tr>
<th>Variable</th>
<th>Low (n=695590 Patients)</th>
<th>Average (n=711141 Patients)</th>
<th>High (n=578267 Patients)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Unadjusted bleeding rate, %</td>
<td>&lt;4.41</td>
<td>4.41–6.23</td>
<td>≥6.24</td>
</tr>
<tr>
<td>Patient characteristics</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Median age, y (IQR)</td>
<td>65.0</td>
<td>65.0</td>
<td>64.0</td>
</tr>
<tr>
<td>Female sex, %</td>
<td>32.7</td>
<td>32.2</td>
<td>32.8</td>
</tr>
<tr>
<td>Non-white race, %</td>
<td>11.2</td>
<td>11.7</td>
<td>13.3</td>
</tr>
<tr>
<td>Median BMI, kg/m² (IQR)</td>
<td>29.1</td>
<td>29.1</td>
<td>29.0</td>
</tr>
<tr>
<td>Medical history, %</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hypertension*</td>
<td>81.8</td>
<td>82.2</td>
<td>81.8</td>
</tr>
<tr>
<td>Diabetes mellitus</td>
<td>36.3</td>
<td>36.4</td>
<td>36.9</td>
</tr>
<tr>
<td>Prior MI</td>
<td>28.9</td>
<td>30.7</td>
<td>31.1</td>
</tr>
<tr>
<td>Prior PCI</td>
<td>41.7</td>
<td>41.1</td>
<td>39.2</td>
</tr>
<tr>
<td>Prior CABG</td>
<td>19.2</td>
<td>19.2</td>
<td>17.5</td>
</tr>
<tr>
<td>Prior CHF</td>
<td>11.1</td>
<td>12.1</td>
<td>12.7</td>
</tr>
<tr>
<td>CVD</td>
<td>12.0</td>
<td>12.5</td>
<td>12.6</td>
</tr>
<tr>
<td>Chronic lung disease</td>
<td>15.0</td>
<td>15.0</td>
<td>15.6</td>
</tr>
<tr>
<td>Current/recent smoker</td>
<td>27.1</td>
<td>27.3</td>
<td>28.5</td>
</tr>
<tr>
<td>Chronic renal failure</td>
<td>2.1</td>
<td>2.4</td>
<td>2.7</td>
</tr>
<tr>
<td>Presentation with STEMI, %</td>
<td>19.8</td>
<td>15.7</td>
<td>13.7</td>
</tr>
<tr>
<td>HF within 2 wk, %</td>
<td>8.3</td>
<td>10.0</td>
<td>11.3</td>
</tr>
<tr>
<td>Cardiogenic shock w/in 24 h, %</td>
<td>2.3</td>
<td>2.7</td>
<td>3.2</td>
</tr>
<tr>
<td>Median preprocedure</td>
<td>13.7</td>
<td>13.7</td>
<td>13.6</td>
</tr>
<tr>
<td>hemoglobin, g/dL (IQR)</td>
<td>(12.4, 14.9)</td>
<td>(12.3, 14.9)</td>
<td>(12.3, 14.9)</td>
</tr>
<tr>
<td>Procedure characteristics</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>PCI status, %</td>
<td>Elective</td>
<td>48.9</td>
<td>42.8</td>
</tr>
<tr>
<td></td>
<td>Urgent</td>
<td>35.9</td>
<td>39.6</td>
</tr>
<tr>
<td></td>
<td>Emergency/salvage</td>
<td>14.9</td>
<td>17.3</td>
</tr>
<tr>
<td></td>
<td>Bifurcation lesion, %</td>
<td>10.6</td>
<td>10.7</td>
</tr>
<tr>
<td></td>
<td>Dissection, %</td>
<td>1.0</td>
<td>1.3</td>
</tr>
<tr>
<td></td>
<td>Perforation, %</td>
<td>0.3</td>
<td>0.4</td>
</tr>
<tr>
<td></td>
<td>Any DES use, %</td>
<td>71.8</td>
<td>71.2</td>
</tr>
<tr>
<td></td>
<td>IABP insertion, %</td>
<td>1.8</td>
<td>2.1</td>
</tr>
</tbody>
</table>

BAS indicates bleeding avoidance strategies; IQR, interquartile range; and PCI, percutaneous coronary intervention.

*Sensitivity Analyses

from this study from <50 to <150 PCIs per site. Using this criterion, 60 sites were excluded. Among the remaining 1236 hospitals, the 5th and 95th percentiles for bleeding rates were 2.6% and 10.3%, respectively, which were similar to the rates of 2.6% and 10.4% from the original analysis.

Figure 3. Correlation between unadjusted and risk-adjusted hospital rankings for post-percutaneous coronary intervention (PCI) Bleeding. Hospitals are ranked by unadjusted bleeding rates on the x axis and by risk-adjusted bleeding rates on the y axis (Spearman correlation coefficient: 0.88).
We also repeated our analyses using the nonrandom effects methodology, which is currently used to generate CathPCI Registry site performance reports. We found similar results with respect to residual variation among hospital post-PCI bleeding rates, even after application of the CathPCI Registry bleeding model. Using fixed effect models, risk-adjustment resulted in a median rank order change of ±86.0 (interquartile range: 35.8, 182.0). After adjustment, there was a reclassification of 25.2% (n=325) of hospitals with respect to outlier status (Table 5). There was 88.8% overall agreement between fixed and random effect strategies to identify hospital outliers (Data Supplement). Post-PCI bleeding was again correlated with hospital use of BAS (Spearman correlation coefficient: −0.26), but not PCI volume or academic status.

### Discussion

Post-PCI bleeding, an important procedural complication associated with poor prognosis, has recently been targeted as a site quality metric, but has not been well characterized among hospitals nationwide. In our analysis of almost 2 million PCI procedures performed at 1292 US hospitals, we observed that hospital bleeding rates varied from 2.1% to 10.3% (5th and 95th percentiles, respectively). From a policy perspective, we found that adjustment for patient clinical characteristics changed hospital outlier classification for >25% of sites and is necessary for appropriate provider comparisons. However, wide variation in hospital bleeding rates persisted after risk-adjustment. Procedural approaches, such as hospital use of BAS, were associated with reduced rates of bleeding, thereby indicating the potential for provider interventions to mitigate PCI bleeding complications.

Assessment of healthcare quality requires an appropriate measure by which to judge provider performance. Consensus criteria for proper selection of performance measures have been previously described. These principles first include the selection of a measure that either represents or is associated with a meaningful outcome to patients and society. Second, the measure must be valid and reliable in its assessment of the process or outcome of interest, and use of this measure to evaluate provider performance must be practically feasible. Finally, consideration of a proper performance measure must account for both patient variability (with the ability to adjust for this variation) and the potential to modify the measure through improvements in processes of care.

Accordingly, post-PCI bleeding may represent an appropriate hospital performance indicator. Bleeding after PCI is associated with increased morbidity, mortality, and cost. Use of bleeding as a performance measure is feasible because of the efficient data collection capabilities of registries such as the CathPCI Registry, and bleeding rates can be adjusted for patient variability through the application of the CathPCI Registry bleeding model, as was performed in our study. Importantly, the last criterion for selection of a suitable performance measure—whether the measure can be modified by changes in processes of care—necessarily depends on (1) the presence of persistent variation after adjustment for patient differences and (2) the availability of strategies to improve performance. In our study, we demonstrated significant residual variation in PCI-related bleeding among hospitals after risk-adjustment, as well as evidence that provider-level alterations in care processes (use of BAS) may reduce bleeding rates.

Several factors may account for the wide variation in hospital-level bleeding after PCI and hospital outliers reported in our study. First, we demonstrate that patient case-mix largely contributes and should be considered when evaluating bleeding rates. Second, differential provider use of BAS may also help to explain differences in post-PCI hospital bleeding. Third, there is likely variability in both ascertainment and reporting of bleeding events. Low bleeding rates may be a result of low site interest in surveillance for bleeding and under-reporting of events. Conversely, sites interested in quality improvement are more likely to report bleeding events and may seem to have higher than average bleeding rates. Although the updated CathPCI Registry bleeding definition tries to address some of these differences in reporting thresholds among sites, accurate assessment of PCI-related bleeding and its use as a performance measure may ultimately require more than participation in a registry and may depend on standardized collection of critical variables at each institution. Finally, some have suggested that the statistical methodology used to create provider reports may be a source of bias. Although there are multiple statistical approaches that can be used to perform site-specific analyses, random effects modeling was chosen for our main analysis because this method uses shrinkage estimators designed to produce estimates that better reflect true hospital effects; furthermore, random effects modeling is becoming widely accepted for profiling of hospital outcomes. We repeated analyses with the nonrandom effects modeling used in the CathPCI Registry site reports, and Risk-adjustment via both approaches resulted in reclassification of outlier status for a large proportion of hospitals (22%–25%). Despite inherent limitations, current CathPCI Registry site performance reports may be useful to incentivize motivated sites to improve practices.
Evaluation of quality improvement strategies from other areas may inform efforts to implement post-PCI bleeding as a hospital performance measure and to reduce bleeding rates among sites. A study by Mehta et al found that implementation of the Guidelines Applied in Practice Initiative, a multifaceted program consisting of caregiver and patient education about key quality indicators, site visits, and guideline-based practice tools, increased adherence to guideline-recommended treatment in the acute myocardial infarction population. Greater adherence to treatment guidelines for acute myocardial infarction patients was also achieved among hospitals routinely using standardized care tools, such as order sets, chart stickers, and discharge checklists. Institutional education and incorporation of BAS and postprocedural bleeding variables into standardized PCI order sets and patient care algorithms might similarly increase awareness of this important quality indicator and help to reduce bleeding rates. Furthermore, although we found an association of hospital BAS use with reduced bleeding after PCI, identification of additional methods for providers to reduce this complication is important and could be achieved through hospital surveys, followed by national dissemination of the most effective approaches. This hospital survey strategy has been successfully used to reduce door-to-balloon times in treating acute myocardial infarction patients.

**Limitations**

Our study has several limitations. First, hospital participation in the CathPCI Registry is voluntary, and self-selected sites may have greater interest in quality improvement, potentially precluding generalization of our results to non-CathPCI Registry–participating hospitals. Second, we could not account for institutional variability in the reporting of postprocedure bleeding, although we tried to reduce the effect of underreporting by excluding sites without any submitted bleeding events. Third, there may be concerns over the inclusion of subjective site-reported variables and blood transfusions in the revised bleeding definition resulting in reporting bias and counting of non-PCI–related transfusions as bleeding events. Nevertheless, the revised bleeding definition does include objective measures of bleeding, and we found consistently higher proportions of all bleeding components at hospitals in the top tertile, as well as similar transfusion thresholds among hospitals across all tertiles. Finally, analyses of hospital factors associated with post-PCI bleeding were hypothesis-generating, and our results should be further investigated. The lack of correlation in our analysis between hospital PCI volume and bleeding rates in contrast to prior studies may be a result of reporting bias, with lower volume sites less likely to report bleeding complications. The modest correlation between hospital BAS use and reduced post-PCI bleeding suggests that other strategies may be important to reduce post-PCI bleeding, but may also be explained by a relatively insensitive hospital-level measure resulting in an underestimation of the true effect of BAS. More accurate assessment of the effect of BAS on bleeding requires further investigation at the patient-level because strategic application of BAS in the highest-risk patients may be the best approach to reduce bleeding.

**Conclusions**

Post-PCI bleeding is associated with adverse patient outcomes and has recently been accepted as a quality of care metric. We demonstrated in a large national data registry that rates of bleeding vary significantly among US hospitals, even after accounting for patient case-mix using a recently revised CathPCI Registry PCI bleeding model. We also found that procedural factors may be important to further reduce bleeding risk. Taken together, our findings support the CathPCI Registry’s use of PCI-related bleeding as a site performance measure and potential incorporation of this metric into other PCI registries. Our results also suggest that provider decisions on procedure methods, such as BAS, may be useful to reduce PCI bleeding. Ultimately, quality improvement initiatives to reduce post-PCI bleeding, perhaps through wide implementation of BAS and sharing of practices from best-performing sites, might lead to improved PCI outcomes, although further investigation is needed.

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**References**


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Supplemental Material

Technical Appendix

We calculated observed bleeding rates for hospitals by dividing the observed number of bleeds by the total number of admissions. The expected number of bleeds for each hospital was determined using the validated CathPCI Registry bleeding model to tabulate the sum of the predicted probabilities of bleeding for each patient at that hospital. Hospital adjusted rates were then obtained by multiplying the ratio of observed to expected number of events by the population bleeding rate. Hospital outlier status for observed rates was defined using the 95% confidence interval (CI) for a hospital’s observed bleeding rate divided by the population bleeding rate: hospitals for which the lower 95% CI limit was greater than one were considered to have high outlier status; hospitals for which the upper 95% CI limit was less than one were considered to have low outlier status; and hospitals whose 95% CI included one were considered to be non-outliers. Hospital outlier status for adjusted rate was defined similarly using the 95% CI for a hospital’s ratio of observed-to-expected bleeding rates. This approach to identifying quality outliers is described in Method #1 in Glance, et al. Med Care. 2006 Apr;44(4):311-9 and is the current methodology used to assess participating CathPCI Registry hospital performance in site reports. However, standard errors were based on the binomial distribution, rather than Poisson, as the former approach is currently used in the NCDR CathPCI site reports.

We also repeated analyses using random effects modeling, which is increasingly used to compare hospital outcomes. In these analyses, unadjusted hospital bleeding rates were calculated with hospital included as a random effect variable, whereas adjusted bleeding rates were determined after accounting for all variables in the CathPCI Registry bleeding model in addition to hospital. Models incorporating hospital as a random effect allow for formal statistical testing
of whether any observed variation in outcome is due to differences among hospitals (variance parameter estimates greater than zero and with p-values <0.05) versus simple sampling variation. Hospitals were again classified as outliers using similar criteria as in the main analysis, but this time using the 95% CI around the hospital’s random intercept. In this approach, described in Method #5 in Glance, et al. Med Care. 2006 Apr;44(4):311-9, the random hospital intercept is a shrunken estimator and represents the log odds of bleeding for each hospital. Hospitals with 95% CI for log odds of bleeding that were greater than 0 were considered upper outliers. Hospitals with 95% CI for log odds of bleeding that were less than 0 were considered lower outliers. All other hospitals were considered to be non-outliers.
Table. Comparison of Adjusted Hospital Outlier Status Defined by Non-random Effects versus Random Effects Models

<table>
<thead>
<tr>
<th>Outlier Status: Non-random Effects</th>
<th>Outlier Status: Random-effects</th>
</tr>
</thead>
<tbody>
<tr>
<td>Low outlier (n=349)</td>
<td>Low outlier (n=349)</td>
</tr>
<tr>
<td></td>
<td>260 (100.0%)</td>
</tr>
<tr>
<td>Non-outlier (n=665)</td>
<td>Non-outlier (n=665)</td>
</tr>
<tr>
<td></td>
<td>0 (0.0%)</td>
</tr>
<tr>
<td>High outlier (n=278)</td>
<td>High outlier (n=278)</td>
</tr>
<tr>
<td></td>
<td>0 (0.0%)</td>
</tr>
</tbody>
</table>